

# US Policy Spillover (?)

China's Accession to the WTO and Rising Exports to the EU

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# Research Question and Hypothesis

## Why did China's EU exports rise so fast after WTO entry?

- ▶ doubts about WTO effect on trade (Rose 2004)
- ▶ no obvious change in conventional channels (i.e. tariffs)

Explore new approach: reduced tariff uncertainty

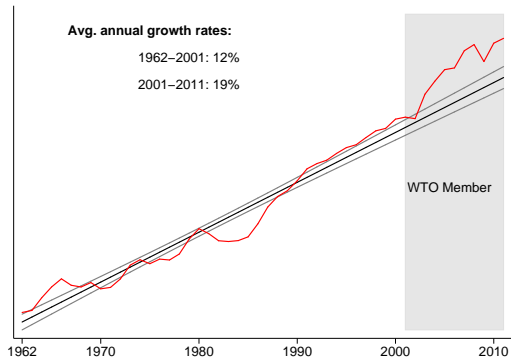
- ▶ China-US trade: more trade after policy change (Handley & Limão 2013)
- ▶ China-EU: no policy change

## Hypothesis

*Spillover of US policies on third countries through economies of scale*

# China's export boom after 2001

## Real Exports to EU-15 countries 1962-2011



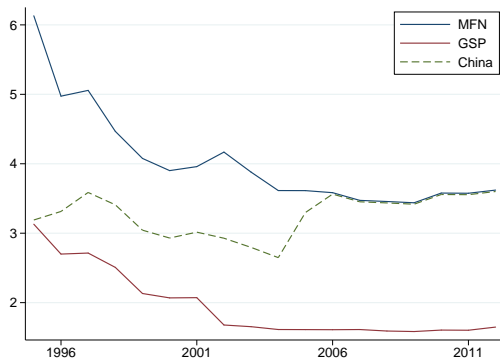
Data: NBER, Comtrade, and PWT 8.0

## WTO Accession (Dec 2001):

- ▶ Tariffs
- ▶ Quotas
- ▶ Other policies

# EU tariffs on Chinese goods

Applicable Tariffs (*ad valorem* equiv) 1995-2012



Data: WITS and EC Regulations (GSP)

WTO Accession (Dec 2001):

- ▶ Tariffs
- ▶ Quotas
- ▶ Other policies

# EU trade policies towards China

## MFA Quotas in Textiles, Clothing, Apparel

- ▶ Dismantling in 2002, 2005, and 2009
- ▶ Decreasing share (1992-2012): 45 → 20%
- ▶ Limited to HS 50-67 (Manuf.: HS 28-96)

## WTO Accession (Dec 2001):

- ▶ Tariffs
- ▶ Quotas
- ▶ Other policies

## Other policies

- ▶ Permanent normal trade relationships (PNTR)
- ▶ MFN/GSP status since 1979/1980
- ▶ No policy change upon WTO access

⇒ Find other source: US policy spillover?

# US Policy Change and Related Studies

## US trade relations with China

- ▶ provisional MFN status since 1980 (Title IV 1974 Trade Act)
- ▶ entailed annual approval by  $\geq 50\%$  votes in US Congress
- ▶ permanent “normal trade relations” (PNTR) since Jan 2002

## The threat of increasing tariffs (before 2002)

- ▶ non-MFN tariffs (“Column-2”)  $\approx 30\%$  higher, on average
- ▶ no abolishment of MFN status but close, esp. in the 1990s
- ▶ e.g. Tiananmen square, NATO bombing, jet accident

## Removal of tariff threat upon WTO entry

- ▶ Handley & Limão 2013: Rising exports to US and lower prices
- ▶ Pierce & Schott 2013: Less US manuf. employment, more trade
- ▶ Feng, Li, Swenson 2014: Increased firm entry, private vs SOEs

# Roadmap

1. Motivation, Background, Literature ✓
2. Theory, Comparative Statics
3. Empirical Analysis/Results
4. Concluding Remarks
5. (Extension: Trade Diversion)

# Theory: Setup

## Develop model that allows for policy spillovers

- ▶ Setup: heterogeneous firms, monopolistic competition (Melitz 2003)
- ▶ Spillover: fixed cost with *bilateral* and *global* component (scale econ.)
- ▶ Tariff uncertainty: expected rate based on two possible scenarios

## Setup: Demand, Supply, and Entry

- ▶ Demand for variety  $j$ :  $x_j = \frac{E_J}{P_J} \left( \frac{p_j}{P_J} \right)^{-\sigma}$
- ▶ Price in destination  $n$ :  $p_{jn} = \left( \frac{\sigma}{\sigma-1} \right) \frac{w}{\varphi_j} d_{Jn} \tau_{Jn}$
- ▶ ZPC productivity:  $\varphi_{Jn}^* = \tau_{Jn}^{\frac{\sigma}{\sigma-1}} \left[ \frac{f_{Jn}}{E_{Jn}(1-\epsilon)} \right]^{\frac{1}{\sigma-1}} \left( \frac{d_{Jn} w}{P_{Jn} \epsilon} \right)$ ;  $\epsilon \equiv (\sigma - 1)/\sigma$

→ More firms export when tariffs and trade costs fall or prices and expend. rise



# Theory: Bilateral Results

Similar to Handley & Limão 2013; Feng et al. 2014

Uncertainty in applied tariffs; two scenarios:  $s = \{p, np\}$

- ▶ Preferential vs. non-preferential tariffs:  $\tau^p \leq \tau^{np}$
- ▶ Probability of switching from  $p$  to  $np$ :  $0 \leq \delta \leq 1$
- ▶ Expected tariff under uncertainty:  $\tau^E = (\tau^{np})^\delta (\tau^p)^{1-\delta}$

Exports to  $n$  under uncertainty in  $n$ :

$$\ln R_{Jn} = -\frac{\sigma k}{\sigma - 1} \ln \tau_{Jn}^E - k \ln d_{Jn} + \frac{k}{\sigma - 1} \ln A_n + \ln \alpha_J - \frac{k - \sigma + 1}{\sigma - 1} \ln f_{Jn}$$

Note:  $A_n$  and  $\alpha_J$  summarize country- and product-specific variables and parameters of the model, respectively.

$$\ln \tau_{Jn}^E = \ln \tau_{Jn}^p + \delta_n \underbrace{(\ln \tau_{Jn}^{np} - \ln \tau_{Jn}^p)}_{\text{Tariff threat: } GAP_{Jn} \geq 0}$$

**Lemma 1:** *The removal of tariff uncertainty in country  $n$ , i.e.  $\delta_n \rightarrow 0$ , has a positive effect on exports to country  $n$ .*

# Theory: Introducing Global Fixed Costs

**Assumption** (Hanson & Xiang, 2011): additive fixed costs;  $f_{Jn} \equiv f_n + f_J$

Exporters consider all potential destinations:

$$\begin{aligned}\tilde{\pi}_{j1} - f_J &= (1 - \epsilon) \left( \frac{w}{\varphi_j \epsilon} \right)^{1-\sigma} \left( \frac{d_{J1}}{P_{J1}} \right)^{1-\sigma} \tau_{J1}^{-\sigma} E_{J1} - f_1 - f_J \\ + \tilde{\pi}_{j2} &= (1 - \epsilon) \left( \frac{w}{\varphi_j \epsilon} \right)^{1-\sigma} \left( \frac{d_{J2}}{P_{J2}} \right)^{1-\sigma} \tau_{J2}^{-\sigma} E_{J2} - f_2 \\ &\vdots \\ + \tilde{\pi}_{jN} &= (1 - \epsilon) \left( \frac{w}{\varphi_j \epsilon} \right)^{1-\sigma} \left( \frac{d_{JN}}{P_{JN}} \right)^{1-\sigma} \tau_{JN}^{-\sigma} E_{JN} - f_N \\ \Leftrightarrow \Pi_j &= (1 - \epsilon) \left( \frac{w}{\varphi_j \epsilon} \right)^{1-\sigma} \sum_{n=1}^N \left[ \left( \frac{d_{Jn}}{P_{Jn}} \right)^{1-\sigma} \tau_{Jn}^{-\sigma} E_{Jn} \right] - \sum_{n=1}^{N^*} f_n - f_J\end{aligned}$$

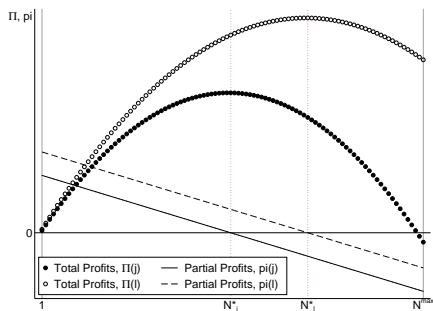
Setting  $\Pi = 0$  gives the multilateral ZPC productivity

$$\Phi_J^* = \sigma^{\frac{1}{\sigma-1}} \left( \frac{w}{\epsilon} \right) \left[ \sum_{n=1}^N [\tau_{Jn}]^{\frac{\sigma}{\sigma-1}} \frac{d_{Jn}}{P_{Jn}} \left( \frac{f_n + f_J}{E_{Jn}} \right)^{\frac{1}{\sigma-1}} \right]$$

# Theory: Implications of Global Fixed Costs

**Lemma 2:** Irrespective of global fixed costs  $f_J$ , a firm  $j$  exports to a destination  $n$  only if bilateral partial profits are positive,  $\tilde{\pi}_{jn} \geq 0$ .

## Partial and Aggregate Profits of two Firms



## Ranking of Destinations:

▶  $\tilde{\pi}_{j1} \geq \dots \tilde{\pi}_{jn} \geq \dots \tilde{\pi}_{jN}$

## Optimal # of Destinations:

▶  $\varphi(j) < \varphi(l) \Rightarrow N^*(j) < N^*(l)$

## Global Fixed Cost Component:

▶ Large  $f_J$  or low  $\varphi(j)$ :  
 $\Pi(N = 1 \dots n) < 0$

**Lemma 3:** If  $\tilde{\pi}_{j1} \geq \dots \tilde{\pi}_{jn} \geq \dots \tilde{\pi}_{jN}$ , and if global fixed costs can be covered, a firm exports to all destinations for which  $\tilde{\pi}_{jn} \geq 0$ .

**Lemma 4:** If  $N = N^*$  is the optimal number of destinations served by any firm  $j$ , then the productivity threshold  $\Phi^*$  increases with  $N$ .

# Theory: Bilateral Tariff Uncertainty

**Numerical example I:** Removal of tariff uncertainty and ZPC thresholds  $\Phi^*$ ; two symmetric countries, set  $\sigma = 3$

Baseline scenario:

- ▶ Tariff uncertainty country 1
- ▶  $\tau_{nJ}^E = (\tau_{nJ}^{np})^\delta (\tau_{nJ}^p)^{1-\delta}$
- ▶  $\tau_{1J}^{np} = 2$ ;  $\tau_{1J}^p = 1$ ;  $\tau_{1J}^E \approx 1.4$

Computed ZPCs and Tariff Uncertainty

	(1)	(2)	(3)
	$\Phi_N^*$	$\Phi_1^*$	$\Phi_2^*$
Baseline: $\tau_1^E = 1.4$	3.53	5.35	3.18
Treatment: $\tau_1^E = 1$	2.90	3.18	3.18

Uncertainty vs. removed uncertainty

- ⇒ First, lowest  $\Phi^*$  in column (3); then in column (1)
- ⇒ Additional firms export to both countries

# Theory: Bilateral Tariff Uncertainty

**Numerical example II:** Reduction of ZPC and size of the policy making country 1

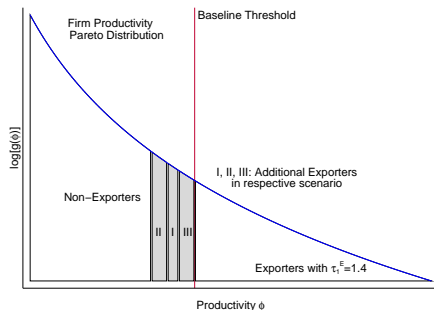
Compute  $\bar{\Phi}$  for  $N = n$

- ▶ asymmetric countries,  $\sigma = 3$
- ▶ tariff uncertainty in  $n = 1$
- ▶  $\tau_{nJ}^{np} = 2$ ;  $\tau_{nJ}^p = 1$ ;  $\tau_{nJ}^E \approx 1.4$

Scenarios: size of countries

- I  $E_1 = 1$ ;  $E_2 = 2$ ;  $E_3 = 0.5$
- II  $E_1 = 2$ ;  $E_2 = 2$ ;  $E_3 = 0.5$
- III  $E_1 = 0.5$ ;  $E_2 = 2$ ;  $E_3 = 0.5$

Computed ZPCs and Tariff Uncertainty



- ⇒ Larger countries have larger effect on multilateral threshold
- ⇒ Large countries absorb larger portion of global fixed cost burden

# Theory: Predictions & Recap

**Predictions:** Exports to  $n$  with global fixed costs

$$\ln R_{Jn} = -\frac{\sigma k}{\sigma - 1} \ln \tau_{Jn}^E - k \ln d_{Jn} + \frac{k}{\sigma - 1} \ln A_n + \ln \alpha_J - \frac{k - \sigma + 1}{\sigma - 1} \ln (f_n + \theta_{Jn} f_J)$$

The parameter  $\theta$  captures the fraction of  $f_J$  covered by  $n$

**Proposition 1:** *A removal of tariff uncertainty in country  $l \neq n$ , increases exports to country  $n$  through a reduction of the global fixed cost burden  $\theta_{n,l} f_J$ .*

**Proposition 2:** *The reduction of the global fixed cost burden  $\theta_{n,l} f_J$ ,  $\forall n \neq l$  implies a reduced ZPC,  $\Phi_J^*$ , and thereby an adjustment at the extensive margin.*

**How realistic is the global fixed cost component?**

- ▶ Hanson&Xiang (2011): strong evidence for services (i.e. US movies)
- ▶ Iacovone&Javorcik (2012): Mexican firms upgrade before exporting
- ▶ Amiti&Freund (2010): China's export growth driven by processing trade
- ▶ If Chinese firms export labor services  $\rightarrow$  decision is not destination-specific

# Empirical Analysis

## Empirical strategy

$$\ln R_{Jn} = -\frac{\sigma k}{\sigma - 1} \ln \tau_{Jn}^E - k \ln d_{Jn} + \frac{k}{\sigma - 1} \ln A_n + \ln \alpha_J - \frac{k - \sigma + 1}{\sigma - 1} \ln (f_n + \theta_{Jn} f_J)$$

Analyze Chinese exports to EU15 after removal of US tariff uncertainty in 2002

- ▶ US tariff uncertainty as of 1999:  $GAP_{J,99} \equiv \ln \tau_{J,US}^{Col2} - \ln \tau_{J,US}^{MFN}$
- ▶ Interaction with period dummy  $D_t^T = 1$  if  $t \geq 2002$

The policy spillover operates through  $\theta$

- ▶ No *change* in uncertainty in the EU:  $\tau_{J,EU}^E = \tau_{J,EU}$
- ▶ Removal of US tariff uncertainty:  $GAP_J^m \times D_t^T \Rightarrow \Delta \theta_{J,EU} f_J$

## Estimation equation

$$\ln R_{Jnt} = b_1 (GAP_J^m \times D_t^T) + b_2 \ln \tau_{Jnt} + b_{Jn} + b_{nt} + b_{St} + \varepsilon_{Jnt} \quad (1)$$

# Data and variables

Main variables and data sources:

- ▶ Chinese exports to the EU: UN Comtrade, HS6 (Rev. 1992), 1995-2005
- ▶ Applied tariffs to China: WITS and EC Regulations of GSP, 1995-2005
- ▶ US Tariff threat: US Tariffs data: NBER, 1988-2001
- ▶ Post WTO-entry dummy:  $D_{t < 2002}^T = 0$ ;  $D_{t \geq 2002}^T = 1$

Summary Statistics: Exports to EU, applied tariffs, tariff threat

Variable		Mean	Std. Dev.	Min	Max
(log) Exports	overall	11.516	2.628	0	22.411
	between		2.473	0	20.696
	within		1.374	1.936	19.040
(log) Tariffs	overall	0.033	0.036	0	0.535
	between		0.035	0	0.325
	within		0.011	-0.087	0.264
U.S. Tariff GAP ( $t > 2001$ )	overall	0.272	0.137	0	1.048
	between		0.137	0	1.048
	within		0	0.272	0.272



# Results: Level of Chinese Exports I

**Proposition 1:** Level of Chinese exports increase after entry to WTO in products with higher  $GAP$

Industry Range	Full (1)	excl. T&C (2)	Full (3)
Spillover $GAP_{J,99}^{U.S}$	0.647** (0.076)	0.237** (0.091)	0.405** (0.077)
EU Tariff $\ln \tau_{Jt}^{EU}$	-0.386 (0.420)	0.283 (0.433)	-0.052 (0.419)
EU Quota removal I $MFA_{J,02}^{EU}$			0.576** (0.034)
EU Quota removal II $MFA_{J,05}^{EU}$			0.449** (0.050)
Observations	270,767	207,476	270,767
R-squared	0.170	0.176	0.172
Fixed effects	$Jn, nt, St$	$Jn, nt, St$	$Jn, nt, St$

Linear panel regressions, based on Eq (1), using data for years 1995-2005 at HS6-destination level.  
Fixed effects: product-destination ( $Jn$ ), destination-year ( $nt$ ), sector-year ( $St$ ).

- ⇒ Average threatened product exported increase by 10.9 percent relative to non-threatened goods.
- ⇒ Given  $k = 4.854$  (Head et al. 2014); Eq (1) and column (1) imply  $\hat{\sigma} = 3.947$ ; larger in other columns.

# Results: Level of Chinese Exports II

## Non-parametric estimation

G1:  $0 < GAP \leq p[25]$ , G2:  $p[25] < GAP \leq p[75]$ , G3:  $p[75] < GAP \leq p[100]$

Industry Range	Full (1)	excl. T&C (2)	Full (3)
G1: $0 < GAP \leq p[25]$	0.192** (0.074)	0.158** (0.076)	0.212** (0.073)
G2: $p[25] < GAP \leq p[75]$	0.400** (0.071)	0.371** (0.073)	0.387** (0.071)
G3: $p[75] < GAP \leq p[100]$	0.463** (0.072)	0.311** (0.076)	0.393** (0.072)
EU Tariff $\ln \tau_{Jt}^{EU}$	-0.524 (0.421)	0.088 (0.435)	-0.181 (0.420)
EU Quota removal I $MFA_{J,02}^{EU}$			0.576** (0.034)
EU Quota removal II $MFA_{J,05}^{EU}$			0.446** (0.049)
Observations	268,499	205,966	268,499
R-squared	0.171	0.177	0.173
Fixed effects	<i>Jn, nt, St</i>	<i>Jn, nt, St</i>	<i>Jn, nt, St</i>

# Results: Extensive vs. Intensive Margin

**Proposition 2:** Policy spillover increases trade at the extensive margin; i.e. more destinations per product.

	Logistic Regressions		Linear Regressions		
	Odd Ratio	Coeff.	# Destinations	Norm. Growth	Log Growth
	(1)	(2)	(3)	(4)	(5)
Spillover	2.511**	0.921**	1.770**	0.159**	0.094 <sup>a</sup>
$GAP_{J,99}^{U.S}$	(0.201)	(0.080)	(0.275)	(0.034)	(0.048)
EU Tariff	0.406*	-0.902*	0.300	-0.010	-0.327
$\tau_{J,t}^{EU}$	(0.169)	(0.417)	(1.071)	(0.327)	(0.405)
EU Quota removal I	1.157*	0.146**	0.920**	0.090**	0.082**
$MFA_{J,02}^{EU}$	(0.040)	(0.034)	(0.144)	(0.017)	(0.023)
EU Quota removal II	2.115**	0.749**	1.234**	0.691**	0.683**
$MFA_{J,05}^{EU}$	(0.148)	(0.070)	(0.210)	(0.039)	(0.060)
Observations	341, 814		44, 038	284, 134	204, 837
R-squared	0.177		0.364	0.056	0.010
Fixed effects	$Jn, t$		$J, St$	$Jn, nt, St$	$Jn, nt, St$

Alternative specifications using data for years 1995-2005.

⇒ All specifications suggest increased entry: Logit (binary); Destinations per  $J$ ; normalized vs log growth.

⇒ Column (4) Normalized Growth rate:  $g^N \equiv \frac{R_t - R_{t-1}}{0.5(R_t + R_{t-1})}$ ;  $g^N \in [-2, 2]$

# Robustness Check: Redistribution of Global Fixed Costs

Does a rise in  $\theta_{J,US}$  really rise exports to the EU?

- ▶ Let  $\theta_{Jn}$  correspond to avg. share of  $n$  in Chinese exports of  $J$
- ▶ Replace  $GAP_J^{US}$  with  $\Delta \bar{s}_J^{US} = \bar{s}_{J,US}^{post} - \bar{s}_{J,US}^{pre}$

	Baseline	Logit		Linear Regressions		
	Levels	Odd Ratio	Coeff.	# Dest.	Norm. vs.	Log Growth
	(1)	(2)	(3)	(4)	(5)	(6)
US Share	0.789** (0.090)	3.371** (0.293)	1.215** (0.087)	2.073** (0.348)	0.030 (0.037)	0.039 (0.058)
EU Tariff	-0.073 (0.420)	0.429* (0.180)	-0.847* (0.419)	0.403 (1.068)	-0.000 (0.327)	-0.328 (0.406)
EU Quota removal I	0.595** (0.034)	1.222** (0.040)	0.201** (0.033)	1.029** (0.142)	0.106** (0.016)	0.091** (0.023)
EU Quota removal II	0.448** (0.050)	2.102** (0.147)	0.743** (0.070)	1.228** (0.213)	0.694** (0.039)	0.685** (0.060)
Observations	367,870	337, 711		43, 307	281, 203	202, 702
R-squared	0.173	0.183		0.374	0.056	0.010
Fixed effects	$Jn, nt, St$	$Jn, t$		$J, St$	$Jn, nt, St$	$Jn, nt, St$

Alternative specifications using data for years 1995-2005.

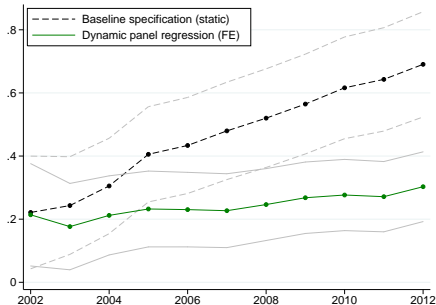
⇒ Goods increasingly exported to the US also grow faster in the EU.

# Robustness Checks: Transition Dynamics

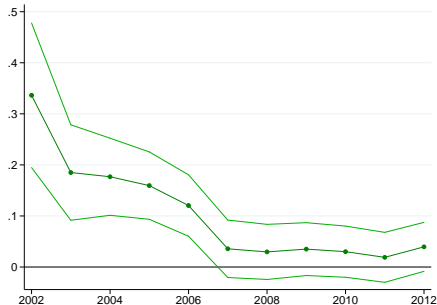
Re-estimating the baseline model with different period length  $T$

Estimated Coefficients  $\hat{b}_1$  for periods 1995- $T$

Level Effect



Growth Effect



Static model biases level effect upwards (dashed line) as period is extended

⇒ Dynamic specification suggests immediate effect around  $\hat{b}_1 = 0.214$ ; implies  $\hat{\sigma} \approx 5$ .

Strongest growth effect in 2002; levels out after few years ⇒ Transitional growth

# Summary & Conclusion

Question: Why did Chinese Exports to EU-15 grow so much faster after 2001?

Motivation: EU policies unchanged, except quota removal in textiles and apparel

Hypothesis: Spillover of US policy change that removed tariff uncertainty for China

Theory: Exporters face global and bilateral fixed costs  $\Rightarrow$  economies of scale

Main Findings:

- ▶ Chinese exports increase more in products where policy change was most felt
- ▶ Increase through export of goods to more destinations (extensive margin)
- ▶ Full effect materializes gradually but within few years after WTO entry

Conclusion:

- ▶ Spillover uncovers important source of trade and international competition
- ▶ Global fixed costs suggest that Chinese firms export labor services

Extensions: *Ad hoc* structural estimates find displacement of US exports by China

**Thank you for your attention!**

# Extension: Trade Diversion?

## Does the policy spillover displace other countries' exports to EU?

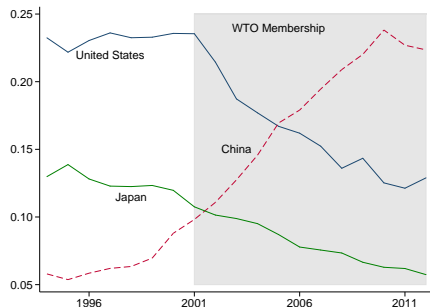
Pierce and Schott (2013)

- ▶ Fall of US manuf employment since 2001
- ▶ Increase of Chinese exports to the US
- ▶ More China-US firm-level transactions
- ▶ Relocation of production?

Fall in US market share and trade diversion?

- ▶ Direct effect of *GAP* on US market share
- ▶ Indirect effect (through increase of China's market share); 2SLS

Market shares in EU-15: 1995-2012



Data: UN Comtrade



# Extension: Reduced Form vs Structural Estimation

Log EU Import Market Share, 1995-2005; Reduced Form vs Structural Estimation

Measure:	Reduced ( $GAP_{J,99}^{U,S}$ )			2SLS ( $\ln Share_{J_t}^{CN}$ )		
	USA (1)	JPN (2)	TUR (3)	USA (4)	JPN (5)	TUR (6)
EU tariff faced $\ln \tau_{J_t}^i$	-1.089 (0.709)	-2.059 <sup>a</sup> (1.143)		-0.945 (0.614)	-0.964 (0.917)	
Policy Change	-0.260* (0.112)	-0.417** (0.158)	-0.003 (0.172)	-0.257* (0.130)	-0.110 (0.178)	-0.499** (0.142)
EU Quota removal I $MFA_{J,02}^{EU}$	0.033 (0.065)	-0.103 (0.084)	-0.186** (0.069)			
EU Quota removal II $MFA_{J,05}^{EU}$	-0.161 (0.099)	0.387** (0.115)	0.007 (0.071)			
EU Tariff (China) $\ln \tau_{J_t}^{CN}$	0.885 (0.635)	1.740 <sup>a</sup> (0.942)	1.110 (0.963)			
Observations	43,749	39,330	32,692	38,402	35,782	30,212
R-squared	0.109	0.038	0.081			
Underidentification (Kleibergen-Paap)				62.66	57.81	70.69
Weak-instruments (Cragg-Donald/KP)				15.92	19.67	17.99
Hansen J-Statistic (p-value)				0.03	0.00	0.21

## US Policy Change

- ▶ Direct negative effects on US and Japan  $\Rightarrow$  Production Relocation?
- ▶ Indirect effects on US and Turkey  $\Rightarrow$  Chinese Competition / Trade Diversion?

# Extension: Chinese Competition at Sector Level

## EU Import Market Shares and Chinese Competition; 2SLS Estimation

Sector	HS Chapter	2SLS estimation ( $\ln Share_{Jt}^{CN}$ )		
		USA (1)	Japan (2)	Turkey (3)
Chemicals	28-38	-0.483** (0.129)	-0.158 (0.172)	0.697 <sup>a</sup> (0.370)
Plastics/Rubbers	39-40	-0.443* (0.178)	-0.097 (0.275)	0.548 (0.639)
Hides/Leather	41-43	-0.545 (0.362)	-1.644* (0.688)	-0.035 (0.519)
Wood Products	44-49	-1.497 <sup>a</sup> (0.777)	0.034 (0.239)	3.678 <sup>a</sup> (1.905)
Textiles	50-60	-0.304** (0.010)	0.036 (0.110)	-0.418** (0.106)
Apparel/Footwear	61-67	-0.542 <sup>a</sup> (0.289)	1.235** (0.434)	-0.364 (0.309)
Stone/Glass	68-71	-0.935 (0.840)	-0.675 (0.969)	-2.209 (2.408)
Metals	72-83	-1.234* (0.010)	-1.186* (0.605)	-0.038 (0.266)
Machinery/Electronic	84-85	-2.255* (1.001)	-1.051 <sup>a</sup> (0.638)	-1.750 (1.572)
Transportation	86-89	-0.818* (0.395)	-0.905 <sup>a</sup> (0.172)	-1.441 (0.915)
Other Manufactures	90-96	-1.287* (0.178)	0.825 (0.559)	-3.770 (3.640)

Statistical significance:  $a = 10\%$ ,  $*$  = 5%,  $** = 1\%$